The most powerful astrophysical events: Gravitational-wave peak luminosity of binary black holes as predicted by numerical relativity

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For a brief moment, a binary black hole (BBH) merger can be the most powerful astrophysical event in the visible Universe. Here we present a model fit for this gravitational-wave peak luminosity of nonprecessing quasicircular BBH systems as a function of the masses and spins of the component black holes, based on numerical relativity (NR) simulations and the hierarchical fitting approach introduced by X. Jiménez-Forteza et al. [Phys. Rev. D 95, 064024 (2017)]. This fit improves over previous results in accuracy and parameter-space coverage and can be used to infer posterior distributions for the peak luminosity of future astrophysical signals like GW150914 and GW151226. The model is calibrated to the $\ell \leq 6$ modes of 378 nonprecessing NR simulations up to mass ratios of 18 and dimensionless spin magnitudes up to 0.995, and includes unequal-spin effects. We also constrain the fit to perturbative numerical results for large mass ratios. Studies of key contributions to the uncertainty in NR peak luminosities, such as (i) mode selection, (ii) finite resolution, (iii) finite extraction radius, and (iv) different methods for converting NR waveforms to luminosity, allow us to use NR simulations from four different codes as a homogeneous calibration set. This study of systematic fits to combined NR and large-mass-ratio data, including higher modes, also paves the way for improved inspiral-merger-ringdown waveform models.

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I. INTRODUCTION

With Advanced LIGO’s [1,2] first detections [3–5], BBH coalescences have become objects of observational astronomy. The peak rate at which BBHs radiate gravitational wave (GW) energy makes them the most luminous known phenomena in the Universe. The source of the first GW detection GW150914 has been inferred to be consistent with two black holes (BHs) of $29^{+4}_{-3} M_{\odot}$ and $36^{+5}_{-4} M_{\odot}$ inspiraling, merging and ringing down as described by general relativity (GR). Its emission of GW energy reached, for a small fraction of a second, a peak rate of $3.6^{+0.5}_{-0.4} \times 10^{56}$ erg/s, equivalent to $200^{+30}_{-20} M_{\odot} c^2 / s$ [3,6,7].

Though this peak luminosity, $L_{\text{peak}}$, is not electromagnetic, but gravitational, we can compare its numerical value to the photon luminosity of other astrophysical sources to illustrate its scale: GW150914 at its peak emitted as much power as $\sim 10^{23}$ suns, $\geq 10^{11}$ times more than all stars in the Milky Way, and still 60–90 times more than the ultraluminous gamma-ray burst GRB 110918A [8].

The peak luminosities for LIGO’s first BBH events were inferred using a fit [9] to data from numerical relativity (NR) simulations, which we will improve upon in this paper through an enhanced fitting method and a significantly larger calibration data set. Source parameters of GW events are determined through Bayesian inference [5,6,10–12].

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1Assuming $L_{\odot} = 3.8 \times 10^{33}$ erg/s, $L_{\text{MW}} = 2 \times 10^{11} L_{\odot}$ and the GRB’s estimated peak isotropic equivalent luminosity of $(4.7 \pm 0.2) \times 10^{51}$ erg/s [8].
comparing LIGO data with waveform models, which are approximate maps between the masses and spins of the binary components and the GW signal. As of the aLIGO O1 run, the state-of-the-art NR-calibrated BBH waveform models were PhenomPv2 [13–15] (a precessing model based on the aligned-spin PhenomD [16,17]) and SEOBNRv2/3 [12,18–21]. A detailed recent study [22] has found these models to be at least sufficiently accurate in the parameter region corresponding to the first detection.

The primary products of this inference are multidimensional sample chains that approximate posterior probability density functions (PDFs) for the intrinsic and extrinsic BBH parameters. Subsequently, such a sampled PDF can be used to infer other quantities, typically obtained through fitting formulas calibrated to NR. Examples include final-state properties [16,23–26]: the final spin and final mass of the merger remnant, a single Kerr BH, which also yield the total radiated energy. In fact, full inspiral-merger-ringdown waveform models include such final-state NR fits to describe the ringdown phase, but due to practical implementation details and for greater flexibility in using updated fits, the values reported in Refs. [3–5,12] come from stand-alone fitting formulas evaluated on posterior PDFs.

The same approach is used for inference of the GW peak luminosity. However, a robust $L_{\text{peak}}$ model for general BBH configurations was not available in the literature prior to O1. An early phenomenological formula [27] is limited to nonspinning BBHs, and thus a new fit [9] had to be developed. To accurately capture the luminosities of the NR calibration set, it also took into account contributions from subdominant harmonics not included in most current waveform models.

In this paper, we present an improved version of that model fit for the GW peak luminosity from the merger of more general BBH systems, including spins on both binary components. Still, we concentrate on cases where the spin of each BH is aligned with the system’s total angular momentum, using the dimensionless components $\chi_i = S_i / m_i^2 = \tilde{S}_i / |\tilde{L}|$ of the spins $\tilde{S}_i$ projected onto the orbital angular momentum $\tilde{L}$. We use the hierarchical data-driven approach introduced in Ref. [26] to develop a three-dimensional Ansatz and fit it to a total of 378 simulations from four separate NR codes, including more subdominant modes than before, and to independent numerical results for large mass ratios obtained with the perturbative scheme of Refs. [28–30]. This addition is essential in producing a well-constrained fit at very unequal masses where NR coverage is sparse or nonexistent.

Notably, the GW peak luminosity $L_{\text{peak}}$ does not depend on the total mass of a BBH system: luminosity generally scales with emitted energy over emission time scale, $L \sim E_{\text{rad}} / \Delta t$. But for a BBH, both the total radiated energy $E_{\text{rad}}$ and the characteristic merger time scale $\Delta t$ are proportional to the total mass, so that $L_{\text{peak}}$ is independent of it. Hence, the GW peak luminosities even of supermassive black hole (SMBH) binaries, observable by eLISA-like missions [31–33] or by pulsar timing arrays (PTAs) [34–36], are similar to those of stellar-mass BBHs. The results of this paper will be applicable to such systems as well.

Besides using $L_{\text{peak}}$ to compare the energetics of GWs and other astrophysical events, one can also consider its relevance for the effect of BBH coalescences on their immediate surroundings. The influence of SMBH mergers on circumbinary accretion disks (see Ref. [37] and references therein) is determined mostly by the integrated radiated energy of the late-inspiral and merger phase, though the authors of Refs. [38,39] suggested weak prompt electromagnetic counterparts sensitive to $L_{\text{peak}}$ and $L_{\text{GW}}(\tilde{t})$. For stellar-mass BBHs, any significant interaction with surrounding material or fields is highly speculative; see e.g. the references in Sec. IV of Ref. [40]. Still it is conceivable that an accurate $L_{\text{peak}}$ model could be useful in constraining exotic models.

Turning $L_{\text{peak}}$ into an independent GW observable through direct signal reconstruction [41–45] will require improved detector sensitivity and calibration accuracy, so that the peak strain can be measured to high absolute accuracy and that degeneracies between the distance estimate and other parameters are significantly reduced. Currently, NR-calibrated fits are the only accurate method to infer peak luminosities from GW observations.

Another motivation for this improved fit is its role as a test case of the general fitting method from Ref. [26] for quantities that require accurate treatment of the higher modes from NR simulations, and of combining NR and perturbative large-mass-ratio results. In these aspects, the present study is a preparation for the development of improved inspiral-merger-ringdown waveform models.

In this paper, we use geometric units with $G = 1, c = 1$ and unit total mass $M$, so that luminosity is a dimensionless quantity, corresponding to $1M$ of energy radiated per $1M$ of time. The conversion factor to watts is $c^5 / G$, and another factor $10^7$ for the usual astronomical unit of erg/s. We will first review the catalog of NR simulations of BBH mergers and perturbative large-mass-ratio data used to construct our model in Sec. II. We discuss the challenges in combining NR data from different simulation codes, the steps taken to process the different sets into a single, effectively homogeneous data set, and how this set is augmented with independent results for large-mass-ratio systems. Given this data set, we discuss the construction and validation of our model fit for peak luminosity in Secs. III and IV. We also compare our new fit with the previous result of Ref. [9], calibrated to a smaller NR data set, that was used during O1 [3–6,12], and to another independent, recently published fit [24]. The Appendix A includes more details on investigations of the NR data.
II. INPUT DATA

A. Numerical relativity data sets

We begin by considering 419 nonprecessing NR simulations from four sources, with their coverage of the three-dimensional BBH parameter space illustrated in Figs. 1 and 2:

(i) 47 simulations performed by the authors with the BAM code [46,47], including those first used in Refs. [16,26].

(ii) 160 simulations from the public SXS catalog [48,49] performed with the Spectral Einstein Code (SpEC) [50].

(iii) 105 simulations from the public GaTech catalog [51,52], performed with the MAYA code [53–56].

(iv) 107 simulations [23,24,57,58] with the LAZEV code [59], labeled “RIT” in the following.

BAM, MAYA and LAZEV are finite difference codes to solve the Baumgarte-Shapiro-Shibata-Nakamura formulation of the GR initial value problem [60] with a singularity-avoiding slicing condition following the moving puncture approach [61,62], whereas the simulations of the SXS Collaboration have been performed with the pseudospectral SpEC code [50] which employs the generalized harmonic gauge (GHG) combined with black-hole excision.

We use mass and spin parameters of the component BHs after equilibration and the initial burst of “junk” radiation. To compute the luminosity for BAM, SXS and GaTech simulations, we begin with the Weyl scalar $\psi_4$ decomposed into its spin-two spherical harmonic multipoles,

$$\psi_{\ell m}(t) = \frac{1}{r} \int_{\Omega} Y_{\ell m}(\theta, \phi) \psi_4(t, \theta, \phi) d\Omega. \quad (1)$$

For SXS $\psi_4$ data, we have applied corrections for center-of-mass drifts [63–65], which remove some unphysical oscillations in higher modes. From these spherical harmonic multipoles, we calculate the GW strain-rate multipoles $\dot{h}_{\ell m}(t)$ via the fixed frequency-integration (FFI) method described in Ref. [66]. We then compute the peak luminosity according to

$$L_{\text{peak}} = \max_{t} \lim_{t \to \infty} \frac{r^2}{16\pi} \sum_{\ell=2}^{\ell_{\text{max}}} \sum_{m=-\ell}^{\ell} |\dot{h}_{\ell m}(t)|^2, \quad (2)$$

truncating the sum over $\ell$ at $\ell_{\text{max}} = 6$. For RIT simulations, we use directly the peak luminosity values as given in Ref. [24], which again include all modes up to $\ell_{\text{max}} = 6$.

NR simulation results always have finite accuracy, and the post-processing from the initial $\psi_4(t, \theta, \phi)$ to the final product $L_{\text{peak}}$ can lead to additional sources of inaccuracies that could substantially affect the accuracy of individual data points and of any NR-calibrated fits. As we aim to fit small subdominant effects, such as those of unequal spins, possible error sources must be carefully analyzed. Thus, we have considered the impact of the following effects on $L_{\text{peak}}$, with details on each aspect given in the Appendix A.
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(i) Conversion between strain $h(t)$, $\psi_4(t)$ and luminosity $L(t)$: The FFI approach is known to be accurate at the 1% level or better. [66] We have also tested its validity by verifying that it agrees with a newly developed alternative technique to compute $h(t)$ and $L(t)$ from the time-domain $\psi_4(t)$.

(ii) Extrapolation effects: Gauge invariance of the waveforms is only well defined for an observer at null infinity. We have extrapolated $L_{\text{peak}}$, as computed from waveforms available at a range of finite extraction radii, to infinity and estimated the extrapolation uncertainty.

(iii) Finite resolution: Convergence tests are only available for a small subset of NR configurations, with estimates indicating that errors due to finite grid resolution are a nondominant but non-negligible contribution to the total error budget.

(iv) Peak accuracy and discreteness: The peak in luminosity can be quite steep, but we have verified that with a sampling of 0.1M in time the difference between two points next to the peak is only on the order of 0.01–0.2%.

(v) Mode selection: While the $\ell = |m| = 2$ mode has the largest peak amplitude for all mass ratios, the importance of other spherical harmonics monotonically increases toward the extreme-mass-ratio limit. We can thus bound higher-mode contributions by comparison with the large-mass-ratio results, where neglecting modes with $\ell > 8$ incurs an error below 1% even for mass ratios of $10^5$ (see below in Sec. II B). For the available NR waveforms, we determine that it is generally sufficient to consider modes up to $\ell_{\text{max}} = 6$, with contributions of the $\ell = 7$, 8 modes rising to only 1% for mass ratio 18 nonspinning cases.

We remove 41 cases from the initial catalog for reasons as discussed in the Appendix A, e.g. because they are inconsistent with equivalent or nearby configurations from the same or other codes. Thus, we perform our fit with a final set of 378 NR results.

B. Perturbative large-mass-ratio data

As the computational cost of NR simulations increases rapidly for unequal mass ratios, no data for BBH systems with mass ratios $q > 18$ are currently available from any of the simulation codes discussed above. ($q = m_1/m_2$ with the convention $m_1 > m_2$.) However, constraining fits to some known behavior in or close to the extreme-mass-ratio limit is essential in ensuring a sane extrapolation towards that limit, and also to reduce uncertainty in the intermediate-mass-ratio region where there is some NR coverage, but it is still very sparse.

For the final-state quantities studied in Ref. [26], we used analytical expressions from Ref. [67] for the limiting case of a test particle orbiting a Kerr black hole. For the peak luminosity, it is known [68,69] that the leading-order term as $\eta \to 0$ must be $L_{\text{peak}} \propto \eta^2$, with the symmetric mass ratio $\eta = (m_1 m_2)/(m_1 + m_2)^2 = q/(1 + q)^2$. However, no fully analytical results for the spin dependence in the extreme-mass-ratio limit exist. Instead, here we constrain our fit by numerical results for finite, but very large mass ratios.

The simulation method for BBH mergers in the test-mass (large-mass-ratio) limit developed in Refs. [28,30,70] combines a semianalytical description of the dynamics with a time-domain numerical approach for computing the multipolar waveform based on BH perturbation theory. The small BH’s dynamics are prescribed using the effective-one-body (EOB) test-mass dynamics, i.e. the conservative (geodesic) motion augmented with a linear-in-$\eta$ radiation reaction expression [28,71]. The latter is built from the factorized and resummed circularized waveform introduced in Ref. [72] (and Ref. [73] for spin) and uses post-Newtonian (PN) information up to 5.5 PN (see also Refs. [74,75] for extensions up to 22 PN). Waveforms are calculated by solving either the Regge-Wheeler-Zerilli (RWZ) $1 + 1$ equations (nonspinning case) or the Teukolsky $2 + 1$ equation (spinning case). Those equations are solved in the time-domain using hyperboloidal coordinates to extract the radiation unambiguously at scri (null infinity) [29,30,76].

The method has been extensively applied for developing and assessing the quality of EOB waveforms [29,30], informing the EOB model in the test-mass limit [77], quasinormal mode excitation [30,70], and computing gravitational recoils [78,79].

The large-mass-ratio waveforms employed here were produced in Refs. [29,30] (RWZ and Teukolsky data, respectively). These waveforms are approximate because $O(\eta)$ effects are not taken into account in the conservative dynamics. However, the consistency of the method was proven in the nonspinning case by showing that, for $\eta \to 0$, the analytical mechanical flux assumed for the small BH’s motion agrees with the numerical GW fluxes to a few percent up to the last stable orbits [29]. The same check has been performed for the spinning case where, instead, significant deviations were found where the spin of the central BH is large ($\chi \gtrsim 0.7$) and aligned with the orbital angular momentum [30]. The discrepancy originates from poor performance of the straight, 5PN-accurate, EOB-resummed analytical multipolar waveform, from which the radiation reaction force is built, for large spins [73]. An iterative method to produce consistent $O(\eta)$ spinning waveforms has been proposed [30], and two such waveforms at $\chi = \pm 0.9$ are available for consistency tests. The method proposed in Ref. [30] is very expensive, since several iterations are needed to find good consistency between the radiation reaction that is used to drive the dynamics and the waveform. The authors of Ref. [75] proposed an additional factorization and resummation of
leaves a three-dimensional parameter space: the mass ratio and two spin parameters $\chi_1$ and $\chi_2$.

As expected, and obvious visually in Fig. 2, one principal direction of curvature in the data set is given by the mass ratio, equivalently expressed as $q$ or the symmetric $\eta$. We can then perform a three-dimensional (3D) hierarchical fit by changing the spin parametrization from the two component spins $\chi_i = S_i/m_i^2$ to a dominant symmetric component

$$\tilde{S} = (m_1^2 \chi_1 + m_2^2 \chi_2)/(m_1^2 + m_2^2)$$

(3)

and a subdominant antisymmetric component $\Delta \chi = \chi_1 - \chi_2$. This is the same choice of effective spin parameter as in Refs. [16,26]. Tests in Appendix C of Ref. [26] show that the fitting method is robust under changing to different parametrizations like the $\chi_{	ext{eff}}$ parameter previously used in Ref. [9].

We perform our fits not on the peak luminosity $L_{\text{peak}}$ itself, but on the rescaled quantity $L'_{\text{peak}} = L_{\text{peak}}/(\eta^2 L_0)$. This removes the expected dominant $\eta^2$ dependence (known analytically for large mass ratios [68,69] and found as the dominant term in previous fits [9,27]) to make small subdominant corrections easier to identify. Here we have also scaled out the equal-mass, zero-spin value $L_0 = L_{\text{peak}}(\eta = 0.25, \chi_1 = \chi_2 = 0)/0.25^2$ to get typical values of order unity. We use $L_0 \approx 0.0164$, the average of the SXS, GaTech and RIT results at this configuration. (These three simulations agree within 0.2%.)

All fits are performed with Mathematica’s NonlinearModelFit function. To avoid overfitting, our model selection is guided by the Akaike and Bayesian information criteria (AICc and BIC, [80,81]), which not only help to choose between fits based on the overall goodness of fit, as measured e.g. by the root-mean-square error (RMSE), but also penalize excessively high numbers of free coefficients. The AICc is defined as

$$\text{AICc} = -2 \ln L_{\text{max}} + 2N_{\text{coeffs}} + \frac{2N_{\text{coeffs}}(N_{\text{coeffs}} + 1)}{N_{\text{data}} - N_{\text{coeffs}} - 1},$$

(4)

with the maximum log-likelihood $L_{\text{max}}$ (assumed Gaussian). This is the standard AIC proposed by Akaike [80] plus a correction for low $N_{\text{data}}$. Schwarz’s alternative criterion [81]

$$\text{BIC} = -2 \ln L_{\text{max}} + N_{\text{coeffs}} \ln(N_{\text{data}}).$$

(5)

despite its name, generally cannot be understood directly as a Bayesian evidence. For specific advantages and disadvantages of these two criteria, their mathematical and philosophical basics and other alternatives see e.g. Ref. [82] and references therein. For both, the model with the lowest value is preferred. The BIC tends to impose a slightly stronger penalty on extra parameters than the AICc,
and we use it as a default ranking of fits, though in practice we do not find disagreements between the two criteria.

### A. One-dimensional nonspinning fit

First, we analyze 84 nonspinning cases, including 81 NR simulations as well as the nonspinning large-mass-ratio data points. As in Ref. [26], we consider several Ansätze choices for the one-dimensional function \( L_{\text{peak}}(\eta, S = 0) \): polynomials up to seventh order, denoted as \( P(m) \), as well as rational functions, denoted as \( R(m, k) \) for polynomial orders \( m \) and \( k \) in the numerator and denominator, respectively. We construct the latter as Padé approximants from an initial polynomial fit to simplify the handling of initial values in the fitting algorithm.

With the dominant \( \eta^2 \) dependence already scaled out, fitting the higher-order corrections allows us to achieve subpercent accuracy, though the additional fit coefficients are not very tightly constrained. The top-ranked fit both by BIC and AICc (with marginally significant differences) is a fifth-order polynomial

\[
L'_{\text{peak}}(\eta, S = 0) = a_5 \eta^5 + a_4 \eta^4 + a_3 \eta^3 + a_2 \eta^2 + a_1 \eta + a_0
\]

with its fit coefficients and their uncertainties given in Table I.

Figure 4 shows this fit, its residuals and comparisons with both the previous fit from Ref. [9] ("T1600018") and the next-highest-ranked alternatives. These next-best alternatives are all rational functions, with the next-simpler polynomial \( P(4) \) disfavored by 7 in BIC and 20% in RMSE and the next-higher-order \( P(6) \) marginally disfavored by 4 in BIC with almost identical RMSE. We find a clear upwards correction over the T1600018 result at low \( \eta \), and differences between high-ranking fits that are much smaller than this correction or the typical residuals. In the data-less region between the lowest-\( \eta \) NR case \( (q = 18) \) and the perturbative results, differences between the highest-ranking fits are larger, but still at most at the same level as the typical fit residuals at higher \( \eta \), corresponding to relative errors below 0.6%. As another comparison, refitting the simple \( L_{\text{peak}}(\eta) = a_2 \eta^2 + a_4 \eta^4 \) Ansatz that we used in Ref. [9] (which in \( L'_{\text{peak}} \) corresponds to just \( \text{const} + \eta^2 \)) is disfavored by over 280 in BIC over this data set, and has a 4 times higher RMSE.

All highly ranked fits agree that the NR data cannot be connected to the large-mass-ratio regime with a simple monotonic function. This behavior might seem surprising, but can be explained by studying the individual modes: the observed behavior of the total peak luminosity results from competing trends of modes that either fall or rise towards \( \eta \rightarrow 0 \). (See Appendix A 6 and Fig. 20 for details, and Refs. [27,83–86] for previous studies of higher-mode amplitudes.) Also we recall that the full \( L_{\text{peak}}(\eta, S = 0) \) is of course monotonic after the dominant \( \eta^2 \) term has been factored back in.

### B. One-dimensional equal-mass-equal-spin fit

Next, we consider 32 equal-mass and equal-spin NR simulations, i.e. configurations with \( \eta = 0.25 \) and \( \chi_1 = \chi_2 \neq 0 \), fitting the one-dimensional function \( L_{\text{peak}}(\eta = 0.25, S) \). We use a similar set of polynomial and rational Ansätze, with the intercept fixed by requiring consistency with the \( \eta \) fit in the nonspinning case, \( L_{'\text{peak}}(\eta = 0.25, S = 0) \). The curvature of this spin dependence at equal masses is relatively mild and can be best fit by a three-coefficient rational function Ansatz.
with the numerical prefactors being due to constructing the Ansatz as a Padé approximant to simplify the handling of Ansätze. In Ref. [26], we constructed 2D final-state Ansätze by

\begin{equation}
L'_{\text{peak}}(\eta = 0.25, \hat{S}) = \frac{0.107b_2\hat{S}^2 + 0.465b_1\hat{S}}{1 - 0.328b_4\hat{S}} + 1.00095,
\end{equation}

with the numerical prefactors being due to constructing the Ansatz as a Padé approximant. This Ansatz is marginally top ranked by both AICc and BIC; it is shown in Fig. 5 and the coefficients $b_i$ are given in Table II. Low-order rational functions are clearly preferred over polynomials, with the P(5) we used in Ref. [9] disfavored by +14 in BIC and having 12% higher RMSE, and the simple R(2,1) Ansatz is sufficient to describe the data to similar subpercent accuracy as the nonspinning set. Adding another term in either the numerator or denominator is possible, but does not improve the statistics; while adding too many terms tends to induce unconstrained coefficients or singularities within the fitting region.

TABLE II. Fit coefficients for the one-dimensional equal-mass-equal-spin $L'_{\text{peak}}(\eta = 0.25, \hat{S})$ fit over 32 data points.

| $b_1$ | 0.9800 | 0.0023 | 0.2 |
| $b_2$ | -0.178 | 0.028 | 15.5 |
| $b_4$ | 1.786 | 0.014 | 0.6 |

C. Spin dependence at large mass ratios

Analyzing the perturbative data at large mass ratios, we verify that the mass-ratio dependence is completely dominated by the leading-order $\eta^2$ scaling in this regime, with the rescaled $L'_{\text{peak}}$ equal to within 0.2% for the three nonspinning data points at mass ratios $q \in \{10^3, 10^4, 10^5\}$.

We treat the single-spin perturbative data as equivalent to results with $\chi_1 = \chi_2$ which is easily accurate enough as $\hat{S}(q = 10^3, \chi_1 = 1.0, \chi_2 = 1.0) - \hat{S}(10^3, 1.0, 0.0) \approx 10^{-6}$ and the spin-difference terms (to be fitted below in Sec. III E) are expected to be suppressed at least by $\eta^2$.

We find that the spin dependence at each of these mass-ratio steps is much steeper than for equal masses, requiring a higher-order spin Ansatz. A spin term at least as complex as R(4,1) is clearly preferred over any lower-order alternatives, with the $q = 10^3$ data yielding $\Delta \text{BIC} \approx 44$ and a factor $>2$ in RMSE in favor of R(4,1) against the R(2,1) found at equal masses, and similar preference even for the two highest mass ratios for which we have only seven data points available each.

Since this analysis of the large-mass-ratio data alone is only used to guide the Ansatz choice in the next section, but not used directly as a constraint, we do not list the detailed results of these fits here. Instead, the final 3D fit using NR and perturbative data will be compared with the high-$q$ data in Sec. IV B.

D. Two-dimensional fits

In proceeding with the hierarchical modeling approach, we can now make a two-dimensional (2D) equal-spin Ansatz informed and constrained by the previous one-dimensional (1D) steps and the large-mass-ratio information. In Ref. [26], we constructed 2D final-state Ansätze by

\[ \text{Here and in the following, we always use “equal-spin” to refer to equal dimensionless spin components } \chi_1 = \chi_2. \]
first simply adding the two one-dimensional fits and then
generalizing each spin coefficient by a polynomial in \( \eta \).
This time, we find that we need to introduce additional
\( \eta \)-dependent higher-order terms in \( \hat{S} \), as the curvature of
\( L'_\text{peak} \) along the spin dimension increases from equal masses
towards the largest mass ratios.

We thus consider a 2D Ansatz of the general form

\[
L'_\text{peak}(\eta, \hat{S}) = L'_\text{peak}(\eta, \hat{S} = 0) + R(m, k)(\eta, \hat{S})
\]

with the \( \eta \) fit from Eq. (6) and the rational function \( R(m, k) \)
in \( \hat{S} \) inheriting the coefficients \( b_i \) from Table II and filled up
with \( b_i = 1.0 \) for orders not present in \( L'_\text{peak}(\eta = 0.25, \hat{S}) \)
from Eq. (7). We then introduce the required freedom to
change the curvature along the \( \eta \) dimension through the substitution

\[
b_i \rightarrow b_i \sum_{j=0}^{J} f_i \eta^j,
\]

with a maximum expansion order \( J \).

On the other hand, the number of free coefficients is
reduced again by consistency constraints with the 1D fits:

\[
f_{i2} = 16 - 16f_{i0} - 4f_{i1} \quad \text{for } b_i \text{ from } \eta = 0.25 \text{ fit},
\]

\[
f_{i2} = -16f_{i0} - 4f_{i1} \quad \text{for other } b_i.
\]

In practice, we use \( R(4,2) \) to match the \( q = 10^3 \) result, thus
introducing one extra power of \( \hat{S} \) in both the numerator and
denominator compared to \( L'_\text{peak}(\eta = 0.25, \hat{S}) \) in Eq. (7).

With 92 equal-spin data points not yet used in the
two one-dimensional subspace fits (including 50 NR
simulations and the single-spin large-mass-ratio results,
which as discussed above can be considered as effectively
equal-spin), we can easily expand the polynomials in \( \eta \) from
Eq. (9) to order \( J = 2 \), \( b_i \rightarrow b_i(f_{i0} + f_{i1}\eta + f_{i2}\eta^2) \), and still
obtain a well-constrained fit. The only further constraint is
that we set the remaining highest-order coefficient in the
denominator, \( f_{i1} \), to zero to avoid a singularity within the
physical \((\eta, \hat{S})\) region, leaving 11 free coefficients.

The resulting fit and its residuals over the equal-spin data
set are plotted in Fig. 7. We again find subpercent relative
errors over most of the calibration set, with an RMSE of
\( 0.0057 \) and only two cases over 1\% relative error (both \( q = 8 \) from BAM).
There is no apparent curvature or oscillatory feature except for
the large-mass-ratio region, where the \( L'_\text{peak} \) quantity plotted in Fig. 7
overemphasizes any remaining features and the corresponding relative errors are
below 0.5\%. This accuracy is similar to that of the large-
mass-ratio-only fits in Sec. III C, thus proving that the
combined two-dimensional fit successfully captures both
the shallow spin slope at similar masses and the steep slope
in the perturbative regime. As discussed in Appendix A7,
several outliers have been removed before the fit; the 2D fit
still matches all equal-spin outliers to below 4\% relative error.

As this equal-spin part of the full \( L'_\text{peak}(\eta, \hat{S}, \Delta \chi) \) will be
refitted, together with unequal-spin corrections, in the next
and final step of the hierarchical procedure, we do not tabulate its best-fit coefficients at this point.

E. Unequal-spin contributions and 3D fit

Simply extending the 2D fit to the full 3D parameter
space either by evaluating fit errors of the equal-spin-only calibrated fit over the whole data set, or by refitting the 2D
Ansatz from Eq. (8), more than doubles the RMSE and
induces oscillations at high \(|\hat{S}|\). But even for such a naive

FIG. 7. Results of the two-dimensional equal-spin \( L'_\text{peak}(\eta, \hat{S}) \) fit. First panel: Comparison of the smooth fit surface with the equal-spin NR data and perturbative results. Second panel: Residuals over the parameter space, color-coded by data provenance.
The next-to-leading-order contributions would be quadratic terms linear in symmetric under relabeling of the component BHs, so that hence, we make the general spin-difference factor, which we therefore expect both in the steps of fixed mass ratio with sufficient numbers of data points. Examples for $q = 1$ and $q = 3$ are shown in Fig. 8. The unequal-spin data set appears more noisy for luminosity than for the final-state quantities studied in Ref. [26], yet can still be analyzed following the same procedure. For each mass ratio step, $q = \{1, 1.33, 1.5, 1.75, 2, 3, 4, 5, 6, 8\}$, we compute the residuals under the nonspinning fit from Eq. (8), then perform four fits in $\Delta \chi$: linear, linear + quadratic, linear + mixed, or the sum of all three terms. Fits of the collected coefficients, as functions of $\eta$, give estimates of the functions $A_i(\eta)$, as displayed with the “per-mass-ratio data” points and “per-mass-ratio fit” lines in Fig. 9.

The scatter of fit coefficients at individual mass-ratio steps is again larger than that found for final spin and radiated energy in Ref. [26], but this procedure still yields sufficient evidence for the existence and shape of a linear spin-difference term and some preference for including both second-order terms, though the data is too noisy to constrain their $\eta$-dependent shape very well. For example, there is an apparent sign switch in the linear term at mass ratio $q = 4 (\eta = 0.16)$, which is most likely due to a combination of the 2D fit being relatively weakly constrained in this region and non-negligible errors in some of the unequal-spin data points, which however cannot easily be discarded as outliers.

The overall fits in $\eta$ are reasonably robust against such problems, and in the next step we will use not this step-by-step analysis, but a more robust fit of the full 3D Ansatz to the full data set, to judge the overall significance of spin-difference terms. A full model selection of $A_i(\eta)$ is clearly not feasible at this point without a more detailed understanding of the point-by-point data quality. Hence, we make very simple choices for the $A_i(\eta)$ with just one power of $\eta$ each,

$$A_1(\eta) = d_{10} \sqrt{1 - 4\eta^2},$$
$$A_2(\eta) = d_{20} \eta^2,$$
$$A_3(\eta) = d_{30} \sqrt{1 - 4\eta^2},$$

and a mixed term proportional to $\tilde{S}\Delta \chi$.

At equal masses ($\eta = 0.25$) BBH configurations are symmetric under relabeling of the component BHs, so that terms linear in $\Delta \chi$ must vanish; this is ensured by the $\sqrt{1 - 4\eta}$ factor, which we therefore expect both in the linear and the mixture term, but not in the quadratic term. Hence, we make the general spin-difference Ansatz

$$\Delta L_{\text{peak}}'(\eta, \tilde{S}, \Delta \chi) = A_1(\eta) \Delta \chi + A_2(\eta) \Delta \chi^2 + A_3(\eta) \tilde{S} \Delta \chi,$$

and a mixed term proportional to $\tilde{S}\Delta \chi$.

We choose the correction terms $\Delta L_{\text{peak}}'$ with guidance from (i) PN analytical results and (ii) an analysis of the residuals of unequal-spin NR simulations under the 2D equal-spin fit. Though PN cannot be expected to be quantitatively accurate for the late-inspiral and merger stages of BBH coalescence—where the peak luminosity emanates from—it can still give some intuition on the qualitative shape of spin and spin-difference effects. The PN spin-orbit flux terms as given in Eq. (3.13) of Ref. [87] and Eq. (4.9) of Ref. [88] include linear terms in $\Delta \chi$ with an $\eta$-dependent prefactor that can be expressed as $P(\eta) / \sqrt{1 - 4\eta}$ with a polynomial $P(\eta)$. The next-to-leading-order contributions would be quadratic in $\Delta \chi$ and a mixed term proportional to $\tilde{S}\Delta \chi$.

FIG. 8. Examples of spin-difference behavior at fixed mass ratios, for scaled NR data $L_{\text{peak}}'$ after subtraction of the 2D ($\eta, \tilde{S}$) fit. First panel: $q = 1$ with the linear term vanishing due to symmetry and mainly quadratic dependence; points in gray are mirror duplicates exploiting the $\chi_1 \leftrightarrow \chi_2$ symmetry at equal masses. There is residual scatter in the $|\Delta \chi| \lesssim 1$ range not captured by the quadratic fit, which is however not much larger than the scatter in equal-spin residuals, and hence probably related to the general uncertainties in NR data quality for $L_{\text{peak}}$. Second panel: $q = 3$ where the linear term dominates and the apparent quadratic dependence likely is noise dominated.
and investigate how much improvement this can yield over the 2D fit.

We now use the full data set except for the 1D subspaces (307 data points, including 265 NR simulations) to fit the full 3D Ansatz from Eq. (11), with the equal-spin and spin-difference contributions from Eqs. (8) and (12)+(13), respectively. The sets of coefficients $a_i$, $b_i$ and $f_{ij}$ are already fixed from the 1D fits and consistency constraints [see Tables I and II and Eq. (10)], leaving between 11 and 14 free coefficients in this final 3D stage. When including all three spin-difference terms, the full Ansatz [with the constraints from Eq. (10) for the $f_{ij}$ still to be applied] is

$$
L'_{\text{peak}}(\eta, S, \Delta \chi) = a_0 \eta^5 + a_1 \eta^4 + a_2 \eta^3 + a_3 \eta^2 + a_4 \eta + a_5 \\
+ \frac{0.107b_2 S^2(f_{22} \eta^2 + f_{21} \eta + f_{20}) + 0.465b_1 \tilde{S}(f_{12} \eta^2 + f_{11} \eta + f_{10}) + \tilde{S}^2(f_{42} \eta^2 + f_{41} \eta + f_{40}) + \tilde{S}^3(f_{52} \eta^2 + f_{51} \eta + f_{50}) - 0.328b_4 \hat{S}(f_{62} \eta^2 + f_{61} \eta + f_{60}) + \hat{S}^2(f_{72} \eta^2 + f_{70}) + 1.0}{
+ d_{20} \eta^3(\chi_1 - \chi_2)^2 + d_{10} \sqrt{1 - 4 \eta^3(\chi_1 - \chi_2)} + d_{30} \sqrt{1 - 4 \eta^3(\chi_1 - \chi_2)}.}
$$

(14)
TABLE III. Summary statistics for the various steps of the hierarchical fit. Note that it is not meaningful to compare AICc and BIC between data sets of different sizes. There is preference for the 3D fit including all three linear + mixture + quadratic terms, although many different choices of the $A_i(\eta)$ Ansätze yield similar results with just ± a few percent in RMSE and ± a few in AICc/BIC, so that the shape of these terms is not yet strongly constrained.

<table>
<thead>
<tr>
<th>$N_{\text{data}}$</th>
<th>$N_{\text{coeff}}$</th>
<th>RMSE</th>
<th>AICc</th>
<th>BIC</th>
</tr>
</thead>
<tbody>
<tr>
<td>1D $\eta$</td>
<td>84</td>
<td>6</td>
<td>$8.13 \times 10^{-3}$</td>
<td>$-817.1$</td>
</tr>
<tr>
<td>1D $\tilde{S}$</td>
<td>32</td>
<td>3</td>
<td>$2.42 \times 10^{-3}$</td>
<td>$-285.8$</td>
</tr>
<tr>
<td>2D($x_1,x_2$)</td>
<td>92</td>
<td>11</td>
<td>$5.65 \times 10^{-3}$</td>
<td>$-751.7$</td>
</tr>
<tr>
<td>2D all</td>
<td>307</td>
<td>11</td>
<td>$1.67 \times 10^{-2}$</td>
<td>$-1914.2$</td>
</tr>
<tr>
<td>3D lin</td>
<td>307</td>
<td>12</td>
<td>$1.51 \times 10^{-2}$</td>
<td>$-2008.0$</td>
</tr>
<tr>
<td>3D lin+quad</td>
<td>307</td>
<td>13</td>
<td>$1.39 \times 10^{-2}$</td>
<td>$-2134.2$</td>
</tr>
<tr>
<td>3D lin+mix</td>
<td>307</td>
<td>13</td>
<td>$1.41 \times 10^{-2}$</td>
<td>$-2082.6$</td>
</tr>
<tr>
<td>3D lin+quad + mix</td>
<td>307</td>
<td>14</td>
<td>$1.36 \times 10^{-2}$</td>
<td>$-2157.8$</td>
</tr>
</tbody>
</table>

We consider residuals and information criteria, summarized in Table III, to check which spin-difference terms are actually supported by the data. These rankings depend on the specific choice of $A_i(\eta)$, but with the current parameter-space coverage and understanding of NR data quality, the main goal is to find general evidence for spin-difference effects and a general idea of their shape, not to exactly characterize them. With the choices made in Eq. (13), we find a 14-coefficient fit with linear-quadratic-mixture corrections that has well-constrained coefficients (see Table IV), is evidently preferred in terms of AICc and BIC, and reduces overall residuals by about 20% in RMSE. Different choices for the powers of $\eta$ in Eq. (13) yield compatible results, while polynomials in $\eta$ with several free coefficients tend to produce underconstrained fits.

TABLE IV. Fit coefficients for the final 3D fit stage, cf. Eq. (14).

<table>
<thead>
<tr>
<th>$B_i$</th>
<th>Estimate</th>
<th>Standard error</th>
<th>Relative error [%]</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>3.79</td>
<td>0.28</td>
<td>7.5</td>
</tr>
<tr>
<td>2</td>
<td>0.402</td>
<td>0.044</td>
<td>10.9</td>
</tr>
<tr>
<td>3</td>
<td>4.27</td>
<td>0.84</td>
<td>19.7</td>
</tr>
<tr>
<td>4</td>
<td>1.62</td>
<td>0.012</td>
<td>0.7</td>
</tr>
<tr>
<td>5</td>
<td>-3.63</td>
<td>0.23</td>
<td>6.3</td>
</tr>
<tr>
<td>6</td>
<td>31.7</td>
<td>1.3</td>
<td>4.2</td>
</tr>
<tr>
<td>7</td>
<td>-274</td>
<td>29</td>
<td>10.4</td>
</tr>
<tr>
<td>8</td>
<td>-0.235</td>
<td>0.011</td>
<td>4.7</td>
</tr>
<tr>
<td>9</td>
<td>6.96</td>
<td>0.44</td>
<td>6.3</td>
</tr>
<tr>
<td>10</td>
<td>0.211</td>
<td>0.022</td>
<td>10.6</td>
</tr>
<tr>
<td>11</td>
<td>1.53</td>
<td>0.45</td>
<td>29.6</td>
</tr>
<tr>
<td>12</td>
<td>3.090</td>
<td>0.044</td>
<td>1.4</td>
</tr>
<tr>
<td>13</td>
<td>-16.7</td>
<td>1.7</td>
<td>10.0</td>
</tr>
<tr>
<td>14</td>
<td>0.836</td>
<td>0.023</td>
<td>2.8</td>
</tr>
</tbody>
</table>

IV. FIT ASSESSMENT

In this section, we assess in some detail the properties and statistical quality of the new three-dimensional peak luminosity fit, with the actual non-rescaled luminosity (in geometric units of $G=c=M=1$) obtained as $\eta^2 L_0 L'_\text{peak}(\eta, \tilde{S}, \Delta \chi)$.

We compare with our previous fit [9] used for LIGO parameter estimation during O1 [3–6,12], which used a much smaller calibration set of 89 BAM and SXS simulations, only modes up to $\ell_{\text{max}} = 4$ and no extreme-mass-ratio constraints; and with the recent Healy and Lousto fit [24] based on 107 RIT simulations, using modes up to $\ell_{\text{max}} = 6$. We attempt to present a fair comparison by analyzing NR and perturbative large-mass-ratio results separately, and also consider the improvement from refitting the unmodified Ansätze of Refs. [9,24] to the present NR data set.

A. Residuals and information criteria

In Fig. 10 we show the distribution of residuals for the 3D fit in $L'_\text{peak}$ projected to the $(\eta, \tilde{S})$ parameter space, so that it can be compared to the 2D results in Fig. 7. The strongest visible outliers in this scaling are at low $\eta$ and correspond to mild actual deviations of at most a 7% relative error at $q = 18$, with 417 of the 423 data points below 3% relative error.

For a comparison with the two previous fits, we first concentrate on the 378 NR simulations only and revisit large mass ratios in Sec. IVB. In Fig. 11 we show histograms of the residuals in $L'_\text{peak}$ for the three fits over this data set, demonstrating that the new fit achieves a narrower distribution. As listed in Table V, the standard deviation of residuals is only half of that for our previous fit.

![FIG. 10. Residuals of the final 3D peak luminosity fit, for NR and perturbative large-mass-ratio data, projected to the 2D parameter space of $\eta$ and $\tilde{S}$. The data sets are distinguished by colors, and unequal-spin points are highlighted with stars.](image-url)
and 3 times lower than for the RIT fit. With a mean offset by only a ninth of a standard deviation, there is no evidence for bias, though that notion is notoriously ambiguous for a data set that samples the parameter space nonuniformly.

The same table contains AICc and BIC values evaluated over the same NR-only data set, which both find a very significant preference for the new fit. Note that, being computed over a different data selection and for \( L_{\text{peak}} \) instead of \( L'_{\text{peak}} \), these values are not directly comparable with the previous Table III. Since we have removed 41 NR cases from the full available data set (see the Appendix A7), it is advisable to check that the statistical preference still holds when including these in the evaluation set. Indeed, the reduction in standard deviations of residuals is then less against the T1600018 and RIT fits, but still roughly 20% and 30%, and there is still a preference of several hundreds in both information criteria.

We also show results for refitting the T1600018 and RIT Ansätze to the present NR + perturbative data set, with the

<table>
<thead>
<tr>
<th>( N_{\text{coef}} )</th>
<th>mean</th>
<th>stdev</th>
<th>AICc</th>
<th>BIC</th>
</tr>
</thead>
<tbody>
<tr>
<td>T1600018</td>
<td>11</td>
<td>3.0×10^{-7}</td>
<td>1.0×10^{-5}</td>
<td>-7732.1</td>
</tr>
<tr>
<td>(refit)</td>
<td>11</td>
<td>-1.8×10^{-6}</td>
<td>4.0×10^{-5}</td>
<td>-6706.0</td>
</tr>
<tr>
<td>HL2016</td>
<td>19</td>
<td>6.9×10^{-6}</td>
<td>1.5×10^{-5}</td>
<td>-7225.5</td>
</tr>
<tr>
<td>(refit)</td>
<td>19</td>
<td>-4.9×10^{-7}</td>
<td>1.0×10^{-5}</td>
<td>-7708.3</td>
</tr>
<tr>
<td>this work</td>
<td>23</td>
<td>-9.8×10^{-7}</td>
<td>4.8×10^{-6}</td>
<td>-8298.1</td>
</tr>
<tr>
<td>(refit)</td>
<td>23</td>
<td>-5.5×10^{-7}</td>
<td>4.8×10^{-6}</td>
<td>-8323.6</td>
</tr>
</tbody>
</table>

FIG. 12. Full NR-calibrated fits from this work and from Refs. [9,24] evaluated at large mass ratios, compared with the same perturbative data (circles, stars and diamonds for mass ratios \( q = \{10^3, 10^4, 10^5\} \)) as in Fig. 6. The T160018 and RIT fits are essentially converged at \( q = 10^3 \) (e.g. 0.4% change at \( S = 1.0 \) for the RIT fit going to \( q = 10^4 \)), and the visually identical lines for higher \( q \) are not shown; our new fit still matches the data at higher \( q \).
spin magnitudes, either negative or positive, and does not reproduce the increasing steepness for even higher mass ratios.

The clearest difference between this fit and the previous ones in the NR-dominated region is for high aligned spins, which is shown in Fig. 13 for the extremal-spin limit, $\chi_1 = \chi_2 = S = 1$, where there is no data available. Both panels give functions of mass ratio $\eta$, and we again compare with the fits from Refs. [9,24]. Top panel: In terms of physical peak luminosity $L_{\text{peak}}$. Lower panel: In terms of rescaled $L'_{\text{peak}} = L_{\text{peak}}/\eta^2 L_0$.

V. CONCLUSIONS

Using the hierarchical analysis approach to the three-dimensional parameter space of nonprecessing quasicircular binary black hole (BBH) coalescences introduced in Ref. [26], we have developed a new model for the peak of the gravitational-wave luminosity of BBH coalescence events, $L_{\text{peak}}$. This model fit is based on the largest-yet combined set of NR results from four independent simulation codes, as well as on perturbative numerical data for the large-mass-ratio regime not currently probed by NR.

The result that BBHs are, for a brief moment during their merger, the most powerful astrophysical events is already clear from dimensional analysis and simplified order-of-magnitude estimates, as is the rough scaling of this peak luminosity with mass ratio. Yet, only detailed NR-calibrated fits allow for a precise understanding of the parameter-space dependence of $L_{\text{peak}}$. Our new fit significantly reduces the residuals for most available NR cases in comparison with a previous version of this fitting procedure [9] used in LIGO O1 data analysis and an alternative fit [24], both calibrated to much smaller data sets.

We also characterized the quality of the luminosity data set considering various sources of NR inaccuracies and the compatibility between different simulation codes, finding that the peak luminosity’s subdominant parameter dependencies are of a similar or even smaller order than typical discrepancies between simulations. This limits the level of detail to which we can model spin-difference effects, though we can still improve over an equal-spin-only fit ($\chi_1 = \chi_2$) and find that the spin-difference dependence qualitatively matches expectations. These statistical improvements, wider parameter-space coverage and systematic understanding of sources of uncertainty can make the new fit a useful ingredient for future parameter estimation studies of BBH events.

The final fit Ansatz is given in Eq. (14), with coefficient estimates listed in Tables I, II and IV. Example implementations of this fit for MATHEMATICA and PYTHON are available as Supplementary Material [90], along with an ASCII table of the data set. The PYTHON implementation is equivalent to that included in the free software LALINFERENCE [91] package.

As more GW detections are made, there will be more opportunities to infer the luminosities of stellar-mass BBH systems. In particular, the aLIGO observatories in the USA [1,2], Advanced Virgo [92] in Italy, and forthcoming observatories in Japan [93,94] and India [95] are poised to become pivotal tools for earth-based GW astronomy, eventually enabling daily BBH detections [96] over a wide parameter range.

The accuracy of the fit presented in this paper should be sufficient for the expected sensitivity at least during the second aLIGO observing run and for “vanilla” BBH events (similar masses, low spins, no strong precession), with

\[ L_{\text{peak}} \sim L_{\text{peak},0} \propto \left| h_{22}(t) \right|^2, \text{ cf. Eq. (2), and } \left| h_{22}(t) \right| \text{ goes to 0 linearly with } \eta \to 0, \text{ so the dominant } L_{\text{peak}} \text{ dependence is } \eta^2 [27]. \]
sampling uncertainties in mass ratio and spins still dominating over fit errors. Still, a continued expansion of the NR calibration set and an improved understanding of higher-mode contributions, precession and the transition from similar-mass to extreme-mass-ratio regimes will be important to improve the understanding of BBH peak luminosities and waveforms.

Meanwhile, this project of fitting peak luminosities is an important step in extending the “Phenom” waveform family [13–17], as our analysis of higher-mode contributions and the demonstration of joint calibration to NR and perturbative large-mass-ratio data can form the basis for improved modeling of full inspiral-merger-ringdown waveforms.

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APPENDIX A: NR DATA INVESTIGATIONS

As a first estimate of the overall accuracy of the peak luminosity data set, we study the differences between results from different codes for equal initial parameters. We then give additional details on the possible error sources listed in Sec. II A and on the properties of higher modes, and discuss the 41 cases not used in the calibration set.

1. Comparison between different codes

To analyze typical deviations between results from different NR codes, we identify simulations with initial BH parameters equal to within numerical accuracy, with a tolerance criterion

\[
|\lambda_i - \lambda_j| \leq \epsilon = 0.0002 \quad \text{for} \quad \lambda_i = \{\eta_i, \chi_{1i}, \chi_{2i}\}. \quad (A1)
\]

This threshold was found in Appendix A of Ref. [26] to be strict enough to reliably identify equivalent initial configurations, and is also tolerant enough to accommodate the minor relaxation of parameters after the initial “junk” radiation which may be different between codes. In Fig. 14 we show the relative difference in \(L_{\text{peak}}\) between such matching cases, including the nonspinning \(q = 4\) case where we have results from all four codes and a few triple coincidences. The set of these tuples is too sparse for clear conclusions on the parameter-space dependence of discrepancies between codes, though there might be some indication of increasing differences at large positive spins, which are particularly challenging to simulate due to increased resolution requirements for capturing the larger metric gradients in the near-horizon zone. We find many pairs with differences below 1%, but also several up to a few % even at not particularly challenging configurations.

This study gives a useful overall estimate of the possible error magnitude on the NR data set: while certainly many simulations are accurate to more than the few-% level, in general for any given simulation that does not have a paired case from another code, or at least nearby neighbors in parameter space, we cannot confidently assume that the errors will be low. This affects in particular the unequal-spin cases, where due to the much larger 3D parameter space very few duplicates exist. On the other hand, for equal spins—and particularly for the densely covered nonspinning or
equal-mass subsets—we can use the duplicates analysis to make a very strict selection of calibration points, allowing the subpercent calibration demonstrated in Secs. III A and III B. The specific decisions are detailed below in Appendix A 7.

As shown in the histograms of Fig. 15, the overall distribution of (relative) differences between equivalent configurations is of a similar width than that of the fit residuals. This demonstrates that we are indeed not over-fitting the data, but also that one would need to characterize the accuracy of all NR cases to a significantly lower level to extract more information on subdominant effects.

2. Luminosity computation from $\psi_4$

Equations (1) and (2) describe the general computation of peak luminosities from the Weyl curvature component $\psi_4(t, \vec{r})$. This conversion is normally performed by either integrating $\psi_4(t, \vec{r})$ twice in time or by first applying a Fourier transform to the data, both to finally obtain the noise present when operating on finite segments of data.

This problem was already solved in Ref. [66] by means of the FFI algorithm, which we briefly describe here. One takes the Fourier-domain strain as

$$\tilde{h}(f) = \begin{cases} \frac{\tilde{\psi}_4(f)}{f}, & \text{if } f \geq f_{\max}, \\ \frac{\tilde{\psi}_4(f)}{f_0}, & \text{if } f < f_{\max}. \end{cases}$$

All physical frequency content must be contained in $[f_0, f_{\text{QNM}}]$ where $f_0$ must be tuned close to the lowest physical frequency for a given mode and $f_{\text{QNM}}$ is the quasinormal mode frequency of the same mode. Thus, a proper selection of $f_0$ down-weights contributions from the low-frequency regime, driving these effects to zero; see Fig. 16.

As a consistency check, we have developed an alternative conversion from $\psi_4(t)$ to $h(t)$ which avoids the step of tuning $f_0$. In Fig. 16 we show an example of the Fourier transform $\tilde{h}_{\text{FFI}}(f)$ of the dominant-mode strain. In general, both local maxima and minima are located in the $[f_{\text{min}} \sim 0.5f_0, f_{\text{max}} \sim 1.2f_0]$ range. The plotted low-frequency behavior occurs for any $\tilde{h}_{\text{FFI}}(f)$ independently of the system’s physical parameters, as a consequence of the finiteness and discreteness of the time-domain waveforms. Empirically we found that the data in $[f_{\text{min}}, f_{\text{max}}]$ can be well fit with an exponential Ansatz, which is then extended to all data in $[0, f_{\text{max}}]$ and combined with the original data above $f_{\text{max}}$:

$$\tilde{h}_{\text{fit}}(f) = \begin{cases} -\frac{\psi_{4,\ell\nu}(f)}{f}, & \text{if } f \geq f_{\text{max}}, \\ a \exp(-(b-f)c) e^{i\phi_{\ell\nu}}, & \text{if } f < f_{\text{max}}, \end{cases}$$

where $\phi_{\ell\nu}$ is the original $\psi_4$ phase. The split in the fit coefficients $a$ (amplitude) and $b$ (peak position) is introduced here so that good starting values for the fit function can be picked more easily. With this approach, we smoothly drive the low-frequency noise to zero, eliminating nonphysical artifacts in the Fourier-domain data.

We find that the difference between peak luminosities from the two different algorithms, when $f_0$ is optimally selected, is generally negligible, e.g. it is about $\sim 0.05\%$ in the example of the $q = 10$ nonspinning SXS waveform, and no significantly larger discrepancies have been found over the data set. So this effect is negligible for our analysis in comparison with other sources of uncertainty.
3. Extrapolation

The NR waveforms used in this paper are extracted at finite radii, which implies ambiguities, in particular due to gauge effects. We therefore extrapolate all waveforms to null infinity, where unambiguous waveforms can be defined. This allows us to assemble a consistent set of peak luminosity values for different codes, and to estimate the errors due to finite radius effects.

However, the extraction properties of the codes are not equal, and thus we have extrapolated the available waveforms following the following prescriptions.

(i) BAM: We have calculated $L_{\text{peak}}$ at each finite radius and then performed a linear-in-$1/R$ extrapolation using only the well-resolved extraction radii. The maximum used for any case is $R \leq 180M$, but for some cases significantly fewer radii can be used for a robust extrapolation, depending on simulation grid resolutions.

(ii) GaTech: $L_{\text{peak}}$ is again calculated at finite radii and then extrapolated with a fit quadratic in $1/R$, only using up to $R \leq 100M$ because the slope generally changes for higher radii; this choice of extrapolation order and radius cut yields the most consistent results with other codes in the analysis of equivalent configurations.

(iii) SXS: These waveforms are already provided at second-, third-, and fourth-order polynomial extrapolation, and we compute $L_{\text{peak}}$ from these data products, after a correction [63–65] for center-of-mass drift, using the second-order extrapolation as the preferred value following Refs. [97,98]. We use waveforms based on the Weyl scalar $\psi_4$, but also compare with waveforms based on a computation of the strain. The SXS $\psi_4$ data use a definition of null-tetrad which is different from their Regge-Wheeler-Zerilli strain data [99–102], and from the definition used in other codes. For the luminosity this difference corresponds to an overall scaling factor of the lapse function to the fourth power as a consequence of the difference between Eqs. (30)–(33) in Ref. [46] and Eqs. (11)–(12) in Ref. [97]. A rough correction for the different tetrad scaling used to compute the Weyl scalar $\psi_4$ is to multiply it by $\alpha^4$ with $\alpha = 1 - 2M_f/R$, where $M_f$ is the final mass and $R$ is an approximation to the luminosity distance using the standard relation with the isotropic radial coordinate for the Schwarzschild spacetime. (Compare also with the analysis in Ref. [103].) Comparisons of SXS luminosities computed from $\psi_4$, strain, and heuristically rescaled $\psi_4$ with data from other codes are included in Figs. 17 and 18.

(iv) RIT: The luminosity data provided in Ref. [24] uses the extrapolation method of Ref. [103].

In Fig. 17 we show the only configuration, the nonspinning $q = 4$ case, for which we have data from all four extrapolation prescriptions.

![FIG. 17. $q = 4$ nonspinning example of extrapolation from finite radii for BAM, SXS and GaTech, with second-order fits for SXS and GaTech and linear for BAM, as well as the RIT value extrapolated with the method of Ref. [103] and its error bar also containing a finite-resolution estimate. In this case we find consistent values from BAM, SXS and RIT, with the GaTech case being an outlier. The $R > 100$ M GaTech data would make the trend more inconsistent, and are excluded from extrapolation. SXS luminosities computed from strain, or from $\psi_4$ but with the $\alpha^4$ rescaling discussed in the text, show a flatter finite-$R$ behavior more similar to the other codes, and extrapolated values consistent with the luminosity from $\psi_4$.]

![FIG. 18. $q = 2.5$ nonspinning example of extrapolation behavior for SXS and GaTech, compared with the extrapolated RIT value. In this case we find consistent GaTech and RIT values, with the asymptotic SXS computed from $\psi_4$ a marginal outlier. Extrapolating the finite-radius peak luminosities computed from $\psi_4$ with rescaling, or from strain, seems to improve consistency in this example, which however needs further study before applying it to the whole data set.]

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codes. This includes peak luminosities computed from the finite-radius strain data available as additional data products from SXS to cross-check the preextrapolated value. We see that extrapolation for \( R \to \infty \) reduces discrepancies in \( L_{\text{peak}} \) between the different codes, but cannot completely alleviate it in this case. Another similar example is shown in Fig. 18 for a \( q = 2.5 \) nonspinning configuration where we have three simulations from SXS, GaTech and RIT, with the GaTech and RIT values being more consistent with each other than with SXS in this case.

The uncertainties of extrapolation fits for BAM, SXS and GaTech can be estimated by the standard deviation on the intersection parameter (equivalent to the confidence interval on the extrapolation to \( 1/R = 0 \)). For the plotted nonspinning \( q = 4 \) case, these are smaller than the remaining largest difference between the results from GaTech and other codes, while for \( q = 2.5 \) the uncertainties are almost wide enough to make the results marginally consistent. For some other cases, these uncertainties can reach up to a few \%, especially when we want to be conservative and take the maximum of (i) the statistical uncertainty for the standard extrapolation-order choice and (ii) the difference between this and the closest alternative order. In general, such an uncertainty estimate cannot provide information about any systematics present in the data from different codes, and indeed for example we find that for BAM the purely statistical extrapolation uncertainties are much smaller in some high-\( q \) cases than for low-\( q \) cases which are generally considered more reliable.

Hence, a study of the extrapolation uncertainties over the whole parameter space is useful in gaining an understanding of the properties of the different codes, but cannot directly be used as a measure of total NR uncertainties.

### 4. Finite resolution

The error contribution from finite numerical resolution can be estimated through convergence tests, reproducing the same configuration at different resolutions. This multiplies computational cost and is hence only practical for a small set of representative simulations. Comparisons of NR results at different resolutions have been discussed e.g. in Refs. [16,47] for the BAM code and in Ref. [104] for GW150914-like SXS and RIT waveforms. For peak luminosities specifically, multiresolution results are available for some BAM and RIT simulations.

The error estimates presented for 107 RIT simulations in Tables XI–XIII of Ref. [24] combine finite-resolution and finite-radius contributions, but for four cases at mass ratios \( q \approx \{1, 1.33, 2.3\} \) and different spins we can extract error estimates due to finite resolution only from Tables XIV–XVII, by comparing \( L_{\text{peak}} \) extrapolated to \( r_{\text{obs}} = \infty \) for the highest finite resolution with \( L_{\text{peak}} \) extrapolated to both infinite radius and infinite resolution. This yields relative error estimates \( \Delta L_{\text{peak}}/L_{\text{peak}} \) of about 0.8–1.6\%.

These estimates fall well within the distributions of our fit residuals and of the “duplicates” study, as shown in Fig. 15, and from comparison with the combined RIT error estimates and with Appendix A3 on extrapolation from finite extraction radius we also see that these two error contributions are typically on a comparable level.

Convergence testing for BAM runs at the particularly challenging \( q = 18 \) mass ratio has previously been discussed in Ref. [16], indicating generally robust behavior. Estimating the finite-difference error as the difference between the highest resolution and a Richardson extrapolation yields < 1\% for both \( \psi_4 \) and \( L_{\text{peak}} \) in the nonspinning \( q = 18 \) case, and for the \( \chi_1 = 0.4 \) simulation we find \( \approx 1\% \) for \( \psi_4 \) and \( \approx 4\% \) for \( L_{\text{peak}} \). We already knew that these simulations at high mass ratios must have wider overall error bars due to e.g. the higher-mode contributions.

Hence, finite resolution can be conjectured to be a nondominant, but also non-negligible contribution to the total uncertainty budget, while a point-by-point evaluation is hindered by the large computational cost.

### 5. Peak accuracy

Since we are dealing with discrete numerical data sets, the peak finding might also be a problem if the sampling is not fine enough, particularly for high mass-ratio cases where the higher modes become more relevant and it is important to sample each mode accurately so that the overall peak profile is not washed out. We have estimated this contribution to NR uncertainties by applying two different time samplings to the data: for the actual \( L_{\text{peak}} \) values used in this paper, we use \( \Delta t = 0.1M \), while here we compare also with a coarser \( \Delta t = 1M \) to illustrate the possible loss of accuracy.

In Fig. 19 we show, for an SXS mass-ratio 10 nonspinning case, that the uncertainty contribution, measured as the difference between two points bracketing the peak, would be about 1\% of the total peak luminosity with the

![FIG. 19. Luminosity computed with different sampling rates in time, for the nonspinning \( q = 10 \) SXS waveform “SXS:BBH:0185” or “BBH_CFMS_d11d2_q10_sA_0_0_0_sB_0_0_0.”](image)
coarser sampling, but is only about 0.05% for the finer sampling that we actually use. As a worst case, we found 0.2% for the nonspinning $q = 18$ BAM result.

6. Mode selection

As introduced in Eq. (2), we compute NR peak luminosities for BAM, SXS and GaTech waveforms as sums over all modes up to $l_{\text{max}} = 6$. The RIT luminosities from Refs. [23,24] use the same cutoff. For the perturbative data from Refs. [28–30] at large mass ratios, we use $l_{\text{max}} = 8$. These choices are based on studying the individual contribution of each mode to the total luminosity, finding that $l > 6$ contributions are sufficiently small to be discarded for the NR data in comparison with other sources of uncertainty.

As an illustrative example, we show in the top panel of Fig. 20 the cumulative peak luminosity when adding modes $\ell$ by $\ell$ (including all $|m| \leq \ell$ at each step) for the $q = 10$ nonspinning SXS waveform, and the per-$\ell$ contributions in the lower panel. The falloff of the higher-$\ell$ contributions to the global peak is expected to be exponential, which is indeed found in this case.

To quantify and extrapolate the loss generally expected for nonspinning configurations, we have estimated the relative loss in $L_{\text{peak}}$ from not including the $\ell = 7, 8$ modes for nonspinning SXS waveforms up to mass ratio $q = 10$ (maximum loss of 0.6%) and the nonspinning BAM simulation at $q = 18$ (loss of 1%), and fit a quadratic function in $\eta$:

$$\frac{\Delta L_{\text{peak}}}{L_{\text{peak}}} = 0.017611 - 0.153760 \eta + 0.334803 \eta^2. \quad (A4)$$

This result is illustrated in Fig. 21, together with a marginally consistent fit when including the $q = 10^3$ Teukolsky result (loss of 2%). The $\ell > 6$ contributions are smaller for

![FIG. 20. Per-mode contributions to the total peak luminosity for the same SXS case as in Fig. 19. Top panel: Cumulative sum up to $\ell$. Lower panel: Natural logarithm of the luminosity contribution per $\ell$. Each point contains all $m$ for the given $\ell$. Similar behavior for large mass ratios was found in Ref. [76].](image)

![FIG. 21. Relative loss in the peak luminosity including modes up to $\ell_{\text{max}} = 6$ against $\ell_{\text{max}} = 8$, for nonspinning SXS cases up to $q = 10$, a nonspinning BAM case with $q = 18$ and the $q = 10^3$ Teukolsky result. Also shown are $q = 18$ and $q = 10^3$ results for $\chi_1 = +0.8$ (above the nonspinning line) and for $\chi_1 = -0.8$ (below), as well as the quadratic nonspinning fit from Eq. (A4) to NR data points only and a fit of the same order including the $q = 10^3$ point, with 90% confidence intervals for both fits.](image)

![FIG. 22. Comparison of rescaled peak luminosities for nonspinning SXS and BAM and perturbative large-mass-ratio data, for a small selection of modes. The points for each mode have been scaled by the maximum for that mode, which is at $\eta = 0.25$ for the 22 mode and at $\eta \to 0$ for the other modes. The connecting lines are fifth-order polynomial fits, which were not statistically optimized and just added to guide the reader’s eye. This can be compared with the sum over modes in Fig. 4. As a guide to the overall strength of the individual modes, we list the nonrescaled maxima $\max_{\eta} \ell_{\text{peak}}(\hat{S} = 0)$ of each of the displayed modes $(\ell m) = \{22, 33, 44, 55, 21, 20\}$.}{0.1}
negative spins and larger for positive spins, as illustrated in the same figure with $\chi_1 = \pm 0.8$ results at $q = 10^3$ and from BAM at $q = 18$. The largest loss for any NR case investigated is $\leq 2\%$ for the $q = 18$, $\chi_1 = +0.8$ BAM case, which is a significant contribution to the overall error budget but still on the level of other error sources. For the perturbative large-mass-ratio results, with a worst-case $\ell > 6$ of $\approx 5\%$, we use $\ell^* = 8$ instead, so that the loss from $\ell > 8$ is limited to $< 1\%$.

Another useful investigation is to consider the $\eta$ dependence, and especially the $\eta \rightarrow 0$ behavior, for individual modes. Fitting $L'_\text{peak}(\eta, \tilde{S} = 0)$ in Sec. III A we found, as illustrated in Fig. 4, that the peak luminosity of all modes summed up to $\ell^* = 6$, after scaling out the dominant $\eta^2$ dependence, is not a monotonic function towards low $\eta$. The increasing relative amplitudes of higher-order modes at low $\eta$ have been studied with NR results previously [27, 83–86], but with our large peak luminosity data set we can now investigate the slope more closely. Repeating the same comparison as in Fig. 4 of rescaled nonspinning peak luminosities between NR (SXS+BAM nonspinning) and
perturbative large-mass-ratio data, but for individual modes, we find—as shown in Fig. 22 for a subset of modes—that these are all monotonic as $\eta \to 0$; however, the slopes are very different, with the dominant 22 mode falling off faster than $\eta^2$ and the subdominant and higher modes falling off much slower, consistent with the general expectation of stronger contributions at low $\eta$. This finding of monotonicity in each mode increases our trust in the combination of NR and perturbative results, and the non-monotonicity of the rescaled peak luminosities after summing the modes can thus be explained as a superposition of these counteracting trends in the individual modes.

7. Cases not used in fit calibration

Of the full catalog of 419 NR simulations from four codes, we have only used 378 to calibrate our new fit. Of the 41 removed cases, 22 are nonspinning or equal-spin configurations. Of these, 17 belong to one of the pairs or groups of equivalent initial parameters identified in Appendix A 1, with differences between the paired results inconsistent at a level higher than the fit residuals we can otherwise achieve in the corresponding subspace fit, or they are individual points inconsistent with an otherwise consistent set of direct neighbors. In these cases we removed from each tuple the case most discrepant with the others and with the global trend. This includes for example the GaTech $q = 4$ and SXS $q = 2.5$ nonspinning cases shown in the extrapolation comparisons of Figs. 17 and 18, or the SXS ($q = 1, \tilde{S} = 0.97$) point whose luminosity seems inconsistent with other $q = 1$, high-spin SXS results.

We emphasize that in the one-dimensional fits for nonspinning and equal-mass-equal-spin BBHs we calibrate the fits to subpercent accuracies, so that this is a very strict criterion for removing cases, which mainly serves to guarantee a very clean calibration of the well-covered subspaces and dominant effects so that in the later steps we have a better chance of isolating and extracting subdominant effects from the general, more noisy data set. In terms of total absolute or relative errors compared with the whole NR data set, several of these cases are not overly inaccurate, and we do not imply that necessarily there are data quality issues with the waveforms from which the luminosities are calculated.

The remaining cases were identified as strong outliers outside of the main distribution in the visual inspection of the two-dimensional equal-spin fit (Sec. III D) and the per-mass-ratio analysis of residuals of unequal-spin cases against the 2D fit (Sec. III E). For these simulations, there are no equivalent or nearby comparison cases, so it cannot be said with certainty whether they would still be outliers in a more densely covered future data set; and at the same time a small residual for any given point is no guarantee for its absolute accuracy when there are no equivalent comparison points. Hence, we have made much less strict exclusions in the sparsely covered unequal-spin range, which limits the accuracy to which we can extract the subdominant spin-difference effects (which are of a similar scale as the remaining scatter in the data set), but also reduces the risk of overfitting to spurious trends in a more strongly trimmed data set.

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